

Friday, Nov 8

Sampling Designs and Inclusion Probabilities

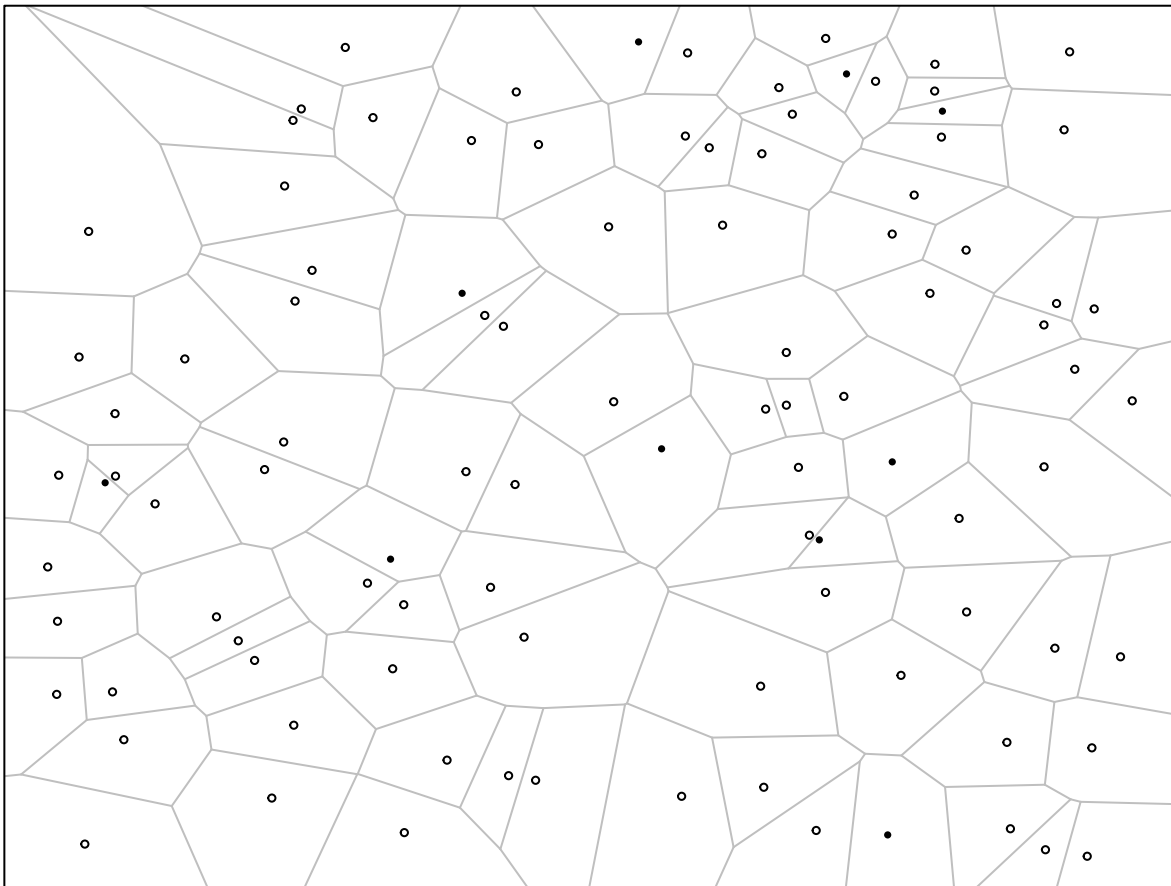
Recall that there are a couple of ways that we can connect the sampling design with the inclusion probabilities.

1. *Indirect approach*: we specify a *sampling design*, which implies the inclusion probabilities of the elements.
2. *Direct approach*: we specify the *inclusion probabilities* and find a sampling design that has those inclusion probabilities.

Indirect Approach Examples

Example: Suppose we have a population of objects distributed over a region with total area A , and we use the following sampling design.

1. Select a random point within the region based on a *uniform distribution* (i.e., select a random x coordinate, and select a random y coordinate, both from uniform distributions).
2. The element that is closest to that point is included within the sample. Do this n times, sampling *with replacement*.



Let y_i be the value of the target variable for the i -th object, and let a_i be the area of the polygon around that object that contains all the points that would lead to the selection of that object. How would we compute an estimate of τ or μ using the Hansen-Hurwitz and Horvitz-Thompson estimators?

Example: Suppose we have a population of elements, and for each element we decide independently with a virtual “coin flip” whether to include it in the sample. This probability could depend on an auxiliary variable x_i so that the probability is cx_i for some $c > 0$ such that $0 < cx_i < 1$ for all elements. This is an example of what is called **Poisson sampling**. A special case of Poisson sampling is **Bernoulli sampling** where all π_i are equal.

What are the first-order and second-order inclusion probabilities? What is the (expected) sample size?

Direct Approach Examples

Example: A variety of sampling *algorithms* have been developed that, given a sampling frame and desired inclusion probabilities, will select a sample using a sampling design with those inclusion probabilities.

Here I will make a small sampling frame for a population of with an auxiliary variable, and compute desired inclusion probabilities as $\pi_i = nx_i/\tau_x$ where $n = 10$.

```
set.seed(123)
library(tidyverse)
f <- data.frame(element = letters[1:25], x = rpois(25, 5)) %>%
  mutate(p = 10 * x / sum(x))
f
```

	element	x	p
1	a	4	0.2721088
2	b	7	0.4761905
3	c	4	0.2721088
4	d	8	0.5442177
5	e	9	0.6122449
6	f	2	0.1360544
7	g	5	0.3401361
8	h	8	0.5442177
9	i	5	0.3401361
10	j	5	0.3401361
11	k	9	0.6122449
12	l	5	0.3401361
13	m	6	0.4081633
14	n	5	0.3401361
15	o	2	0.1360544
16	p	8	0.5442177
17	q	3	0.2040816
18	r	2	0.1360544
19	s	4	0.2721088
20	t	9	0.6122449
21	u	8	0.5442177
22	v	6	0.4081633
23	w	6	0.4081633
24	x	11	0.7482993
25	y	6	0.4081633

The R package **sampling** includes several sampling algorithms for specified inclusion probabilities. Each of them creates an indicator variable to tell us if a given element is to be sampled.

```
library(sampling)
f <- f %>% mutate(I = UPtille(p))
f
```

	element	x	p	I
1	a	4	0.2721088	0
2	b	7	0.4761905	1
3	c	4	0.2721088	0
4	d	8	0.5442177	1
5	e	9	0.6122449	0
6	f	2	0.1360544	0
7	g	5	0.3401361	0
8	h	8	0.5442177	0

```

9      i  5 0.3401361 1
10     j  5 0.3401361 0
11     k  9 0.6122449 0
12     l  5 0.3401361 1
13     m  6 0.4081633 1
14     n  5 0.3401361 1
15     o  2 0.1360544 0
16     p  8 0.5442177 0
17     q  3 0.2040816 1
18     r  2 0.1360544 0
19     s  4 0.2721088 0
20     t  9 0.6122449 0
21     u  8 0.5442177 1
22     v  6 0.4081633 0
23     w  6 0.4081633 1
24     x 11 0.7482993 1
25     y  6 0.4081633 0

```

So here is our sample.

```
f %>% filter(I == 1)
```

```

  element x      p I
1      b  7 0.4761905 1
2      d  8 0.5442177 1
3      i  5 0.3401361 1
4      l  5 0.3401361 1
5      m  6 0.4081633 1
6      n  5 0.3401361 1
7      q  3 0.2040816 1
8      u  8 0.5442177 1
9      w  6 0.4081633 1
10     x 11 0.7482993 1

```

Next step is to go observe the target variable for these elements.

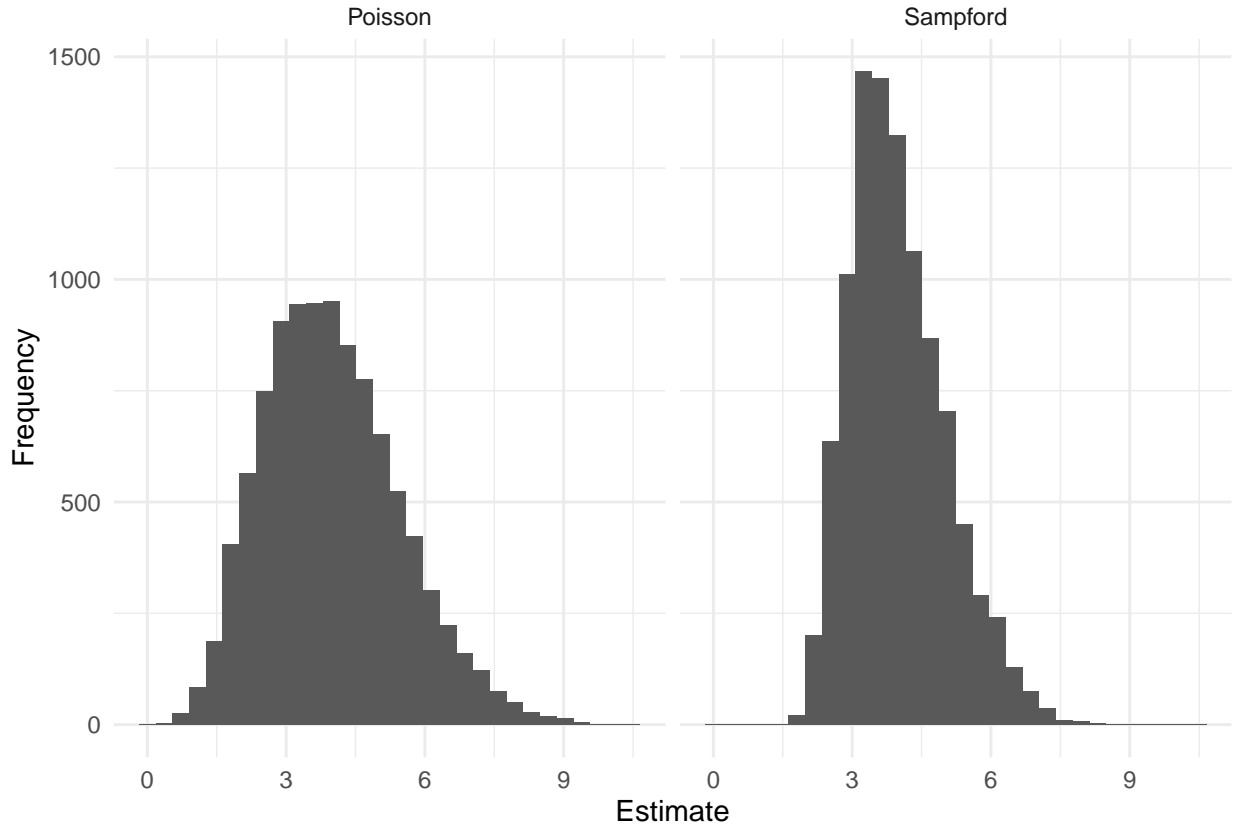
Not all sampling algorithms are equally effective in terms of the *variance* the Horvitz-Thompson estimator. Why?

Here are the results of a simulation study using two different sampling algorithms, Poisson and Sampford, with the sampling frame given above (target variable values were generated randomly).¹

¹Sampford's method is relatively simple. First sample just one unit using the inclusion probabilities $\pi_1, \pi_2, \dots, \pi_n$. This is easy because when $n = 1$ the inclusion probabilities are the same as the selection probabilities. Next sample the remaining $n - 1$ units *with replacement* with selection probabilities are computed as

$$\delta_i = \frac{\pi_i / (1 - n\pi_i)}{\sum_{k=1}^N \pi_k / (1 - n\pi_k)}.$$

But if the sample does not contain *distinct* units (i.e., some elements appear more than once) then repeat both steps until it contains only distinct units. It can be shown that this results in a sample of n units with the specified inclusion probabilities.



design	variance	bound
Poisson	2.175507	2.949920
Sampford	1.072398	2.071133

There are almost always multiple sampling designs with the same inclusion probabilities. Some algorithms (e.g., Sampford) tend to result in estimators with lower variance than others (e.g., Poisson). In some cases we can use additional information in the form of auxiliary variables to “guide” the choice of sampling design.

Balanced Sampling

Suppose that we have the following information:

1. Specified (first-order) inclusion probabilities (π_i) for each element in a population of elements.
2. The values of one or more auxiliary variables ($x_{i1}, x_{i2}, \dots, x_{ik}$) for each element in the population of elements, as well as the their totals ($\tau_{x_1}, \tau_{x_2}, \dots, \tau_{x_k}$).

A **balanced sampling** design has the following properties.

1. The inclusion probabilities of the elements equal the specified values.
2. The sample is (approximately) calibrated such that

$$\tau_{x_1} \approx \sum_{i \in S} \frac{x_{i1}}{\pi_i}, \tau_{x_2} \approx \sum_{i \in S} \frac{x_{i2}}{\pi_i}, \dots, \tau_{x_k} \approx \sum_{i \in S} \frac{x_{ik}}{\pi_i}.$$

These are sometimes called the *balancing equations*.

Whereas ratio and (generalized) regression estimators seek to calibrate the sample by re-weighting *after* the sample is obtained, while balanced sampling seeks to select a sample that is *already calibrated*.

Example: Two estimators with “unbalanced sampling” and one with balanced sampling for the [Red Oaks population](#).

```
library(BalancedSampling)
library(trtools) # ask me for instructions as to how to install this package

set.seed(123) # random number seed

y <- redoak$volume
x <- with(redoak, pi * (dbh/2)^2)

reps <- 10000 # number of samples to sample
N <- nrow(redoak) # population size
n <- 50 # sample size
mux <- mean(x) # population mean of auxiliary variable

est.unb <- rep(NA, reps) # estimates from unbalanced sampling (no auxiliary variable)
est.rat <- rep(NA, reps) # estimates from unbalanced sampling (ratio estimator)
est.bal <- rep(NA, reps) # estimates from balanced sampling

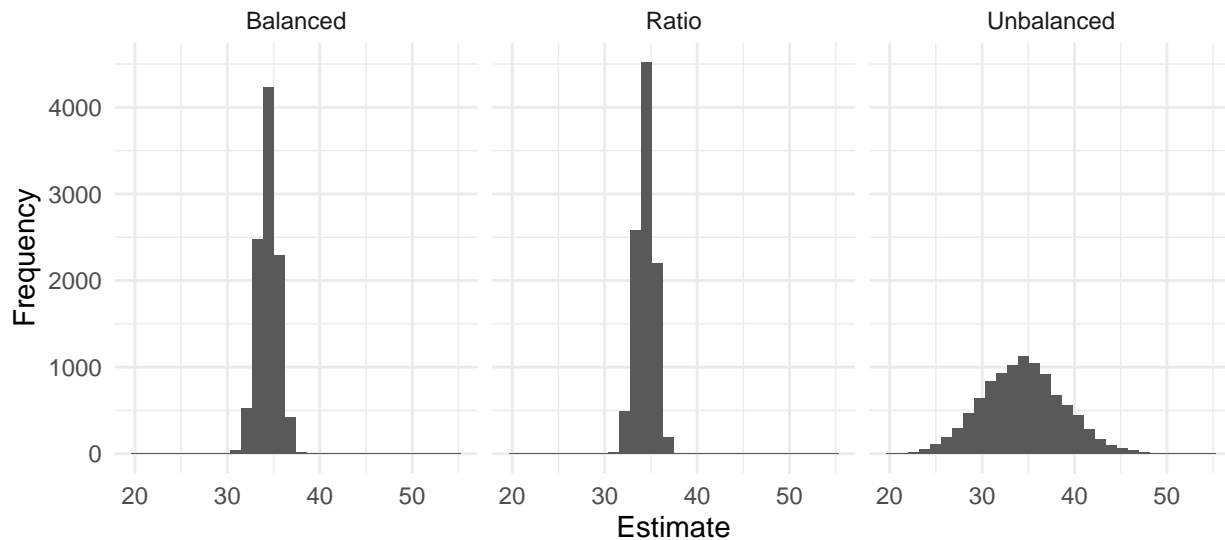
p <- rep(n/N, N) # specify inclusion probabilities
X <- cbind(p, x) # specify array (matrix) of auxiliary variables

for (i in 1:reps) {
  indx <- sample(1:N, n) # unbalanced sampling (SRS)
  est.unb[i] <- sum(y[indx]/p[indx])/N
  est.rat[i] <- mux*sum(y[indx]/p[indx])/sum(x[indx]/p[indx])
  indx <- cubestratified(p, X, rep(1, N)) # balanced sampling using the "cube method"
  est.bal[i] <- sum(y[indx]/p[indx])/N
}

# visualization of the approximated sampling distributions

d <- data.frame(estimate = c(est.unb, est.rat, est.bal),
  design = rep(c("Unbalanced", "Ratio", "Balanced"), each = reps))

library(ggplot2)
p <- ggplot(d, aes(x = estimate)) + theme_minimal() +
  geom_histogram() + facet_wrap(~design) +
  labs(x = "Estimate", y = "Frequency")
plot(p)
```



```
library(dplyr)
d %>% group_by(design) %>% summarize(variance = var(estimate), bound = 2*sqrt(variance))
```

```
# A tibble: 3 x 3
  design      variance bound
  <chr>      <dbl> <dbl>
1 Balanced      1.17  2.16
2 Ratio         0.973 1.97
3 Unbalanced    19.1   8.74
```

Example: Here is a simulation study with a population of 1000 elements. For each element we have two auxiliary variables: x_{i1} and x_{i2} . I used two sampling designs to select a sample of $n = 50$ elements: an *unbalanced* sampling design where all elements have an inclusion probability of 0.05 (SRS), and a *balanced* sampling design using the same inclusion probabilities but also using the two auxiliary variables for the balancing equations.

```
library(BalancedSampling)

set.seed(123) # random number seed

reps <- 10000 # number of samples to sample
N <- 1000    # population size
n <- 50     # sample size

est.bal <- rep(NA, reps) # array for estimates from balanced sampling
est.unb <- rep(NA, reps) # array for estimates from unbalanced sampling

x1 <- runif(N, 0.1, 25) # first auxiliary variable
x2 <- runif(N, 0.1, 25) # second auxiliary variable
y <- rpois(N, x1 + x2) # target variable

p <- rep(n/N, N) # specify inclusion probabilities
X <- cbind(p, x1, x2) # specify array (matrix) of auxiliary variables

for (i in 1:reps) {
  indx <- sample(1:N, n) # "unbalanced" sampling (SRS)
  est.unb[i] <- sum(y[indx]/p[indx])/N # compute estimate of mu
  indx <- cubestratified(p, X, rep(1, N)) # balanced sampling using the "cube method"
```



```

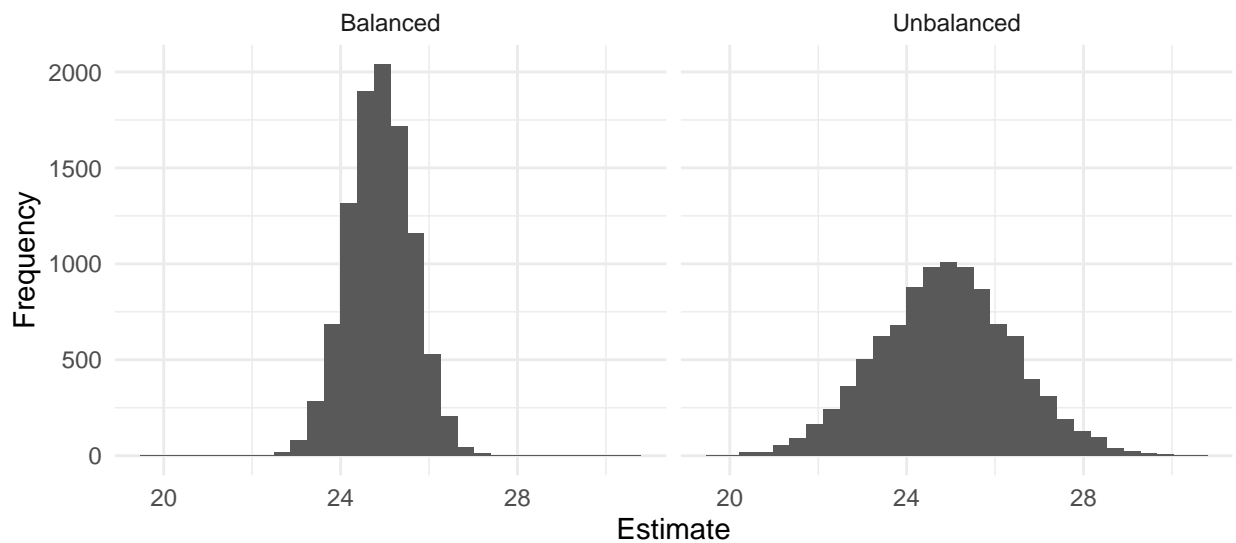
  est.bal[i] <- sum(y[indx]/p[indx])/N    # compute estimate of mu
}

# visualization of the approximated sampling distributions

d <- data.frame(estimate = c(est.unb, est.bal),
  design = rep(c("Unbalanced", "Balanced"), each = reps))

library(ggplot2)
p <- ggplot(d, aes(x = estimate)) + theme_minimal() +
  geom_histogram() + facet_wrap(~design) +
  labs(x = "Estimate", y = "Frequency")
plot(p)

```



```

library(dplyr)
d %>% group_by(design) %>% summarize(variance = var(estimate), bound = 2*sqrt(variance))

```

```

# A tibble: 2 x 3
  design    variance bound
  <chr>      <dbl> <dbl>
1 Balanced    0.510  1.43
2 Unbalanced  2.34   3.06

```

An interesting use of balanced sampling is with strata. Stratified random sampling can be viewed as a special case of balanced sampling, but balanced sampling allows us to make use of overlapping strata.

Balanced Sampling via Inclusion Probabilities

Suppose we obtain a balanced sample with the auxiliary variable $x_i = \pi_i$. What then does the balancing equation imply?

Now suppose we set the inclusion probabilities such that

$$\sum_{i \in \mathcal{D}} \pi_i = n_d,$$

where \mathcal{D} denotes a domain, and we specify an auxiliary variable of

$$x_i = \begin{cases} \pi_i, & \text{if } i \in \mathcal{D}, \\ 0, & \text{otherwise.} \end{cases}$$

What then does the balancing equation imply?

Suppose the inclusion probabilities are defined as $\pi_i = nx_i/\tau_x$. Then we can show that the sample will be balanced with respect to x_i .

Stratified Random Sampling as Balanced Sampling

Suppose we use balanced sampling with the following.

1. Set the inclusion probabilities such that $\pi_i = n_j/N_j$ if the i -th element is in the j -th stratum.
2. Use auxiliary variables $x_{i1}, x_{i2}, \dots, x_{iL}$ such that

$$x_{ij} = \begin{cases} 1, & \text{if the } i\text{-th element is from the } j\text{-th stratum,} \\ 0, & \text{otherwise.} \end{cases}$$

Balanced sampling will satisfy the balancing equations if the number of elements from the j -th stratum is n_j . The balancing equations for x_{ij} are

$$\tau_{x_j} \approx \sum_{i \in \mathcal{S}} \frac{x_{ij}}{\pi_i} = \frac{N_j}{n_j} \sum_{i \in \mathcal{S}} x_{ij}.$$

Since $\tau_{x_j} = \sum_{i=1}^N x_{ij} = N_j$, the balancing equations will be satisfied exactly if $\sum_{i \in \mathcal{S}} x_{ij} = n_j$ (i.e., the number of elements sampled from the j -th stratum is n_j). It turns out this is the same as stratified random sampling, and if the variances of the strata are approximately equal it is approximately an optimal allocation (i.e., proportional allocation).

Balanced Sampling for Overlapping Strata

Balanced sampling gives us a way to deal with *overlapping strata*. For example, suppose we can stratify Hobbits by *Farthing* (North, South, East, or West), and by age group (young, adult, old). Suppose we know the number of elements in each Farthing ($N_n, N_s, N_e,$ and N_w), and the number of elements in each age group ($N_y, N_a,$ and N_o), but *not* the number of elements in each *combination* of Farthing and age group (e.g., we do not know how many young Hobbits there are in the North Farthing). So we cannot use stratified random sampling, but we can use balanced sampling if we specify auxiliary variables for each stratum such as

$$x_{in} = \begin{cases} 1, & \text{if the } i\text{-th element is from the North Farthing,} \\ 0, & \text{otherwise,} \end{cases}$$

for the North Farthing stratum, and

$$x_{iy} = \begin{cases} 1, & \text{if the } i\text{-th element is from young age group,} \\ 0, & \text{otherwise,} \end{cases}$$

for the young age group stratum. Note that we would have to specify auxiliary variables for the other Farthing and age group strata.

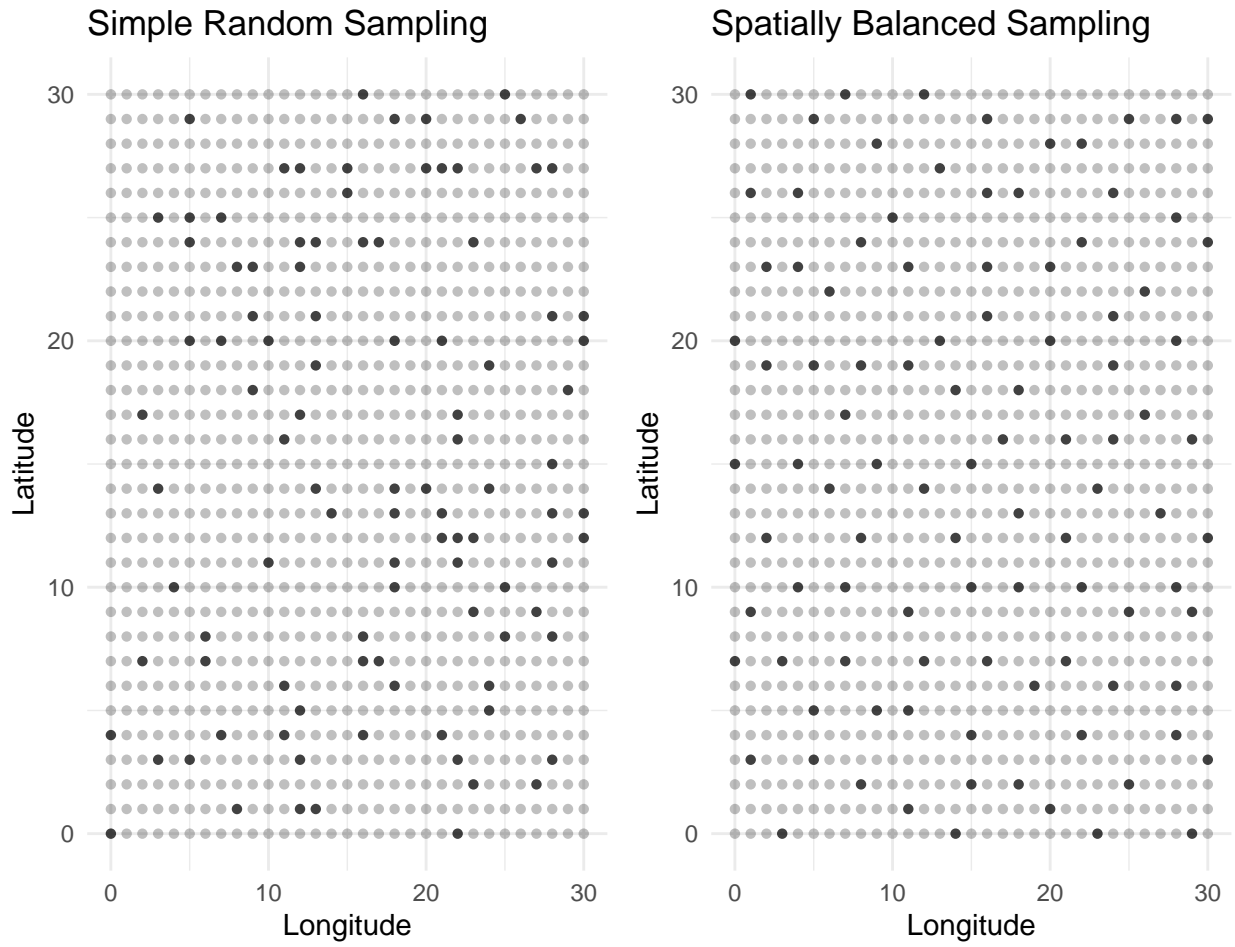
There is a loose connection here with the technique of *raking* which is a way of calibrating the sample *after* it is selected with respect to overlapping strata. Here we use balanced sampling to try to select a balanced sample.

The R package **BalancedSampling** includes functions for balanced sampling.

Spatially Balanced Sampling

A **spatially balanced sampling** design has the following properties.

1. The inclusion probabilities of the elements equal the specified values.
2. The elements in the selected sample tend to “spread out” in terms of the inter-element distances. Here the auxiliary variables effectively reflect *where* the elements are located in space/time.



The R packages **BalancedSampling** and **Spbsampling** include functions for spatially balanced sampling.